

## B5440 – Exercise 3

### Exercises

1. Construct an example of a pair of random variables  $(T, U)$  representing failure times and censoring times that are dependent but the “independent censoring” assumption holds.
2. (Fleming and Harrington Exercise 1.9) Suppose  $T, U$  have joint distribution given by  $H(t, s) = P(T > t, U > s) = \exp(-\lambda t - \mu s - \theta ts)$  where  $0 \leq \theta \leq \lambda\mu$ .
  - a. Find the marginal survival functions and corresponding hazard functions for  $T$  and  $U$ .
  - b. Calculate

$$\alpha_1^\#(t) = \frac{-\frac{\partial}{\partial u} P(T \geq u, U \geq t)|_{u=t}}{P(T \geq t, U \geq t)}$$

and

$$\alpha_2^\#(t) = \frac{-\frac{\partial}{\partial u} P(T \geq t, U \geq u)|_{u=t}}{P(T \geq t, U \geq t)}.$$

- c. Suppose  $\lambda = 1, \mu = 2$  and  $\theta = 2$ . Plot and compare i) the marginal survival function for  $T$  to  $H_1^\#(t) = \exp(-\int_0^t \alpha_1^\#(s) ds)$  and ii) the marginal survival function for  $U$  to  $H_2^\#(t) = \exp(-\int_0^t \alpha_2^\#(s) ds)$ . What do you notice and why is this comparison meaningful?
  - d. Suppose one erroneously assumes  $T$  and  $U$  are independent, and takes their joint distribution to be

$$H^\#(t, s) = H_1^\#(t)H_2^\#(s).$$

Find  $H^\#(t, s)$  and observe that  $H^\#(t, t) = H(t, t)$ .

- e. Suppose the observed data consist of  $\{(\min(T_i, U_i), I(T_i \leq U_i)), i = 1, \dots, n\}$ . The crude survival functions are defined by  $Q_1(t) = P(T > t, U > T)$  and  $Q_2(t) = P(U > t, T > U)$ . Show that whether the joint distribution of  $(T, U)$  is given by  $H(t, s)$  or  $H^\#(t, s)$  as calculated in part d, one obtains the same  $Q_1(t)$  and  $Q_2(t)$ .
3. For a fixed  $t$ , a  $1 - \alpha\%$  confidence interval for the  $S(t)$  the survival probability at  $t$  years based on an estimate  $\hat{S}(t)$  with estimated standard error  $\hat{se}(t)$  is

$$\hat{S}(t) \pm z_{1-\alpha/2} \hat{se}(t),$$

where  $z_{1-\alpha/2}$  is the upper  $\alpha/2$  quantile of a standard normal distribution. This manner of constructing confidence intervals provides the guarantee that for the year  $t$ , the true survival will between the lower and upper limits, over approximately  $(1 - \alpha)\%$  of repeated samples.

If we consider the collection of pointwise confidence intervals in the range  $[a, b]$

$$\{\hat{S}(t) \pm z_{1-\alpha/2}\hat{s}e(t) : a \leq t \leq b\}$$

we have defined a *confidence band* for the survival curve, which is the process or collection  $\{S(t) : a \leq t \leq b\}$ . We have gone from defining an interval on a one-dimensional quantity to defining a continuous banded region for a function of time. We would like a confidence band  $\{[l(t), u(t)] : a \leq t \leq b\}$  that has simultaneous coverage properties, i.e.,

$$P\{l(t) \leq S(t) \leq u(t) : a \leq t \leq b\} \approx 1 - \alpha.$$

Demonstrate or argue why this naive construction of a confidence band based on the collection of pointwise confidence intervals does not satisfy the simultaneous coverage property.

## Reading

ABG Chapter 3, sections 3.1.1-2, 3.1.5-6, 3.2.1, 3.2.4-6, 3.3.1-3,5.

Hall and Wellner 1980 (on canvas).